


BMJ Open Association between chiropractic spinal manipulation and gabapentin prescription in adults with radicular low back pain: retrospective cohort study using US data

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ABSTRACT

Objectives Radicular low back pain (rLBP) is often treated off-label with gabapentin or by chiropractors using chiropractic spinal manipulative therapy (CSMT). To date, no studies have examined the association between these interventions. We hypothesised that adults under 50 years of age receiving CSMT for newly diagnosed rLBP would have reduced odds of receiving a gabapentin prescription over 1 year follow-up.

Design Retrospective cohort study.

Setting US network including linked medical records, medical claims and pharmacy claims of >122 million patients attending large healthcare organisations (TriNetX), queried 15 June 2023, yielding data from 2017 to 2023.

Participants Adults aged 18–49 were included at their first occurrence of rLBP diagnosis. Exclusions were severe pathology, other spinal conditions, on-label gabapentin indications and gabapentin contraindications. Propensity score matching controlled for variables associated with gabapentin use and receipt of prescription medication over the preceding year.

Interventions Patients were divided into CSMT or usual medical care cohorts based on the care received on the index date of rLBP diagnosis.

Primary and secondary outcome measures OR for gabapentin prescription.

Results After propensity matching, there were 1635 patients per cohort (mean age 36.3±8.6 years, 60% women). Gabapentin prescription over 1-year follow-up was significantly lower in the CSMT cohort compared with the usual medical care cohort, with an OR (95% CI) of 0.53 (0.40 to 0.71; $p<0.0001$). Sensitivity analyses revealed early divergence in cumulative incidence of prescription; and no significant between-cohort difference in a negative control outcome (gastrointestinal medication) suggesting adequate control for pharmacological care preference.

Conclusions Our findings suggest that US adults receiving CSMT for newly diagnosed rLBP have significantly reduced odds of receiving a gabapentin prescription over 1-year follow-up compared with those receiving usual medical care. Results may not be generalisable and should be replicated in other healthcare

STRENGTHS AND LIMITATIONS OF THIS STUDY

- ⇒ Study methods were crafted by an interdisciplinary team with the aim of minimising bias.
- ⇒ This study incorporated a new-user design, including patients at the first occurrence of a diagnosis of radicular low back pain, to make cohorts more homogeneous and comparable.
- ⇒ While we controlled for several variables via propensity matching to make cohorts more similar with respect to the likelihood of receiving a gabapentin prescription, variables such as income and pain severity were unavailable or poorly represented in the data set.
- ⇒ Although this study included several thousand patients, it may only be generalisable to large integrated academic healthcare settings in the USA.
- ⇒ Given that this study is observational and may have residual confounding, it should be repeated using a prospective study design.

settings and corroborated by a prospective study to reduce confounding.

BACKGROUND

The USA has the leading age-standardised prevalence of low back pain (LBP) in the world.¹ Together, low back and neck pain account for the leading cause of medical expenditures in the USA.² LBP can be divided into subtypes according to pathophysiology. Radicular low back pain (rLBP), which involves a nerve root lesion, is considered a type of neuropathic pain, and involves radiating symptoms into the ipsilateral lower extremity.^{3,4} Conversely, non-rLBP resulting from myofascial, discogenic, sacroiliac or zygapophyseal joint pain is considered nociceptive and does not necessarily radiate to the lower limb.^{3,4} Consequently, the subtype



of LBP pathophysiology influences its pharmacological treatment approach.⁵

Gabapentin is an anticonvulsant, anti-epileptic medication, used as first-line therapy for several types of neuropathic pain including diabetic neuropathy and herpetic neuralgia.^{6,7} Gabapentin may alleviate neuropathic pain by binding to a subunit of voltage-gated calcium channels, subsequently inhibiting ectopic nerve discharges.^{6,7} Considering this mechanism of action, gabapentin has also been used off-label to treat neuropathic symptoms of LBP, namely rLBP.^{5,7}

While gabapentin has had supporting evidence and US Food and Drug Administration (FDA) approval for use in neuropathic pain conditions since 1993,^{8,9} systematic reviews in 2018 and 2022 demonstrated clear evidence of lack of its effectiveness for rLBP.^{10,11} Additionally, there is growing evidence of its risks including abuse, misuse, dependence and withdrawal.⁹ Potentially deleterious adverse effects of gabapentin include somnolence, dizziness, ataxia and fatigue, as well as new-onset asthenic symptoms, particularly in patients with muscular problems.¹²

Accordingly, several clinical practice guidelines do not recommend gabapentin for the treatment of LBP or rLBP,¹³ including those of the American Family Physician (2017).¹⁴ Evidence supporting the use of gabapentin for LBP is considered inconclusive by guidelines from the North American Spine Society (2020),¹⁵ Global Spine Care Initiative (2020)¹⁶ and Veterans Affairs/Department of Defense (2019 and 2022).^{17,18} Furthermore, gabapentin prescription for LBP has been described as a marker of low-value care¹⁹ and medical overuse.²⁰

Despite the paucity of evidence, and in contrast to clinical guideline recommendations, gabapentin continues to be commonly prescribed for LBP. A survey of 545 US adults (mean age 52 years (range 20–92)) in 2018 revealed that 20% of patients who visited a medical doctor for LBP had been recommended gabapentin in the preceding 12 months.²¹ A cross-sectional study examining over 230 000 outpatient visits in the USA between 2011 and 2015 found that 99% of gabapentin prescriptions were for off-label indications; the most common were degenerative spinal disorders and other back problems, together accounting for 27% of prescriptions.²² In addition, there were increasing rates of episodes of prescription of gabapentin (relative increase of 440%) and concomitant opioid and gabapentin prescription (relative increase of 344%) in the USA between 2006 and 2018.²³

Chiropractors are portal-of-entry providers in the USA who frequently treat spinal disorders.^{24–26} When treating rLBP, these providers often use chiropractic spinal manipulative therapy (CSMT),²⁵ a hands-on treatment directed to the joints of the spine.²⁷ CSMT is supported by systematic reviews^{28,29} and recommended by clinical practice guidelines for the treatment of LBP^{14,15,17} and rLBP.^{30,31}

Although chiropractors cannot prescribe medications such as gabapentin within their scope of practice,²⁴ previous studies have found that the initial type of provider seen for LBP influences the subsequent likelihood of

receiving a prescription for certain medications.^{32–34} These studies have found that patients initiating care for LBP with a chiropractor compared with other providers have reduced odds of receiving an opioid or benzodiazepine prescription.^{33,35,36} However, to our knowledge, no research has explored the association between receipt of chiropractic care versus usual medical care for LBP and the likelihood of subsequent gabapentin prescription.

Considering that gabapentin is commonly prescribed off-label for rLBP, against spine and pain care guideline recommendations, the present study examined if undergoing CSMT influenced the subsequent likelihood of receiving a gabapentin prescription after rLBP diagnosis.

Objectives

This study examined the relationship between CSMT versus usual medical care and subsequent gabapentin prescription among patients newly diagnosed with rLBP identified from a large US database. We hypothesised that adults receiving CSMT on the index date of rLBP diagnosis would have reduced odds of receiving a gabapentin prescription compared with those receiving usual, non-chiropractic medical care over 1-year follow-up.

MATERIALS AND METHODS

Study design

This study incorporated a retrospective observational cohort design using aggregated and linked medical records, medical claims and pharmacy claims data, and implemented new-user, active comparator features to improve cohort comparability and reduce bias.^{37,38} An a priori protocol for the present study was registered in the Open Science Framework in January 2023 (<https://osf.io/rt6f3>).³⁹ Our manuscript reporting adheres to the Strengthening the Reporting of Observational Studies in Epidemiology statement.⁴⁰ Following peer review at BMJ Open, we made three changes to our methods in June 2023, in which we (1) added a cumulative incidence graph to illustrate the timing of gabapentin prescription, (2) propensity matched for receipt of any prescription medication over the year preceding the index date to better account for patients' potential preference to receive pharmacological care⁴¹ and (3) examined for the likelihood of prescription of a negative control outcome medication⁴² (any gastrointestinal medication) over the follow-up year to further explore patients' potential preferences towards pharmacological care, with the latter two changes replacing our previous E-value sensitivity analysis. As practice guidelines and prescribing patterns for gabapentin have evolved over time, only data from the most recent 5-year span were included (15 June 2017 to 15 June 2023). To allow for a 1-year follow-up for included patients, only patients with an index diagnosis of rLBP up to 1-year preceding the query date (15 June 2023) were included (enrolment ending 15 June 2022). To help ensure patients were not lost to follow-up, patients were required to have at least one additional healthcare

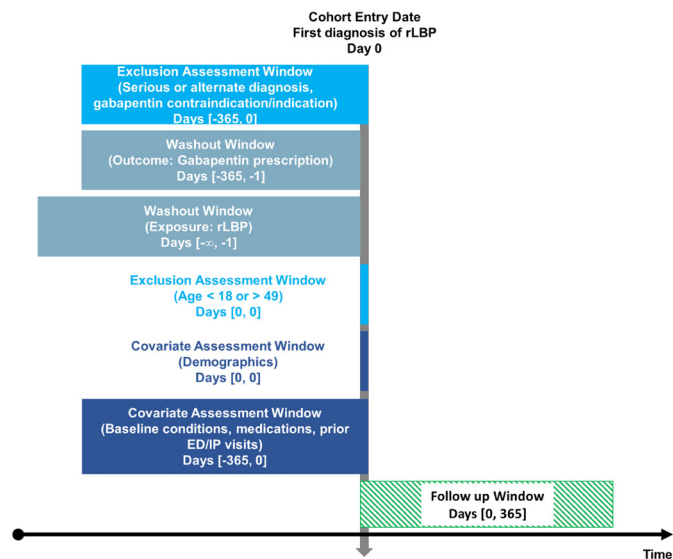


Figure 1 Graphical depiction of study design. The vertical arrow represents the index date of diagnosis of radicular low back pain. Assessment windows to the left of the vertical arrow represent time periods occurring before this date over a span of days (#,#). The ‘∞’ indicates that the time window extends as far previous as data are available per patient. Windows overlapping with the vertical arrow occur on the date of index diagnosis. The follow-up window occurring after the index diagnosis is indicated by a green rectangle. Image created by Robert Trager using creative commons template from Schneeweiss *et al*,⁷⁹ using Microsoft PowerPoint V.2206. ED, emergency department; IP, inpatient; rLBP, radicular low back pain.

encounter of any kind during the year following the index date of rLBP diagnosis (figure 1).

Setting and data source

Study data were sourced from a US research network (TriNetX, Cambridge, Massachusetts, USA),⁴³ which includes aggregated, de-identified data from linked electronic medical records, medical claims and pharmacy claims of 122 million patients. This network includes 84 large academically affiliated healthcare organisations and their outpatient, inpatient and specialty offices, which remain anonymous per data use agreement. The database is searched using standard terminology, such as the International Classification of Disease (ICD) codes. A centralised TriNetX team routinely assesses the data set for completeness, conformance and plausibility.^{43 44} A prior study estimated that medication data was at least 87% complete in the TriNetX data set.⁴⁵ At University Hospitals, access to the TriNetX network is managed by Clinical Research Center personnel.

Data regarding the characteristics of chiropractors in the included study sites also remain anonymous. However, chiropractors in integrated healthcare organisations are typically employed within physical medicine and rehabilitation or physical therapy offices, and have on average 21 years of clinical experience with over 6 years working in the integrated care setting.⁴⁶

Participants

Eligibility criteria

Inclusions

Patients aged 18–49 years were included at the first occurring (index) date of rLBP diagnosis. Only patients with rLBP were included, as this type of LBP often involves neuropathic pain, which is the suggested therapeutic target for gabapentin.⁷ The washout period for rLBP extended as far as data were available preceding the index diagnosis date (which varied per patient), such that patients had no prior recorded diagnosis of rLBP. The current study definition for rLBP included ICD codes that describe sciatica and lumbosacral radiculopathy (online supplemental table 1).⁴⁷ This definition did not include diagnoses related to disc degeneration, disc herniation and spondylosis, which may cause axial LBP without radiculopathy.⁴⁸

Neuropathic pain is more common in those with LBP related to lumbar disc herniation compared with lumbar stenosis, scoliosis or spondylolisthesis.⁴⁹ The age bracket of adults under 50 was selected as rLBP is more likely to result from lumbar disc herniation in patients of this age,^{50–52} while older patients are more likely to have lumbar stenosis underlying rLBP.⁵³ Focusing on a narrower population with rLBP in the current study aimed to create a participant pool with more homogeneous acute pathophysiology, as the likelihood of neuropathic pain (ie, the therapeutic target of gabapentin) varies across LBP aetiologies.⁴⁹

Patients were divided into two cohorts based on receipt of CSMT versus usual medical care. The CSMT cohort served as the test cohort, while the cohort receiving usual medical care served as an active comparator. Patients receiving CSMT on the date of index diagnosis of rLBP were included in the CSMT cohort, while patients not receiving CSMT on the date of index diagnosis formed the cohort receiving usual medical care (online supplemental table 2). In the USA, treatment codes describing CSMT are used almost exclusively by chiropractors.⁵⁴ Usual medical care was defined for the purposes of this study as any of a range of medical services besides CSMT, including physical therapy, medications and interventional or surgical procedures.

Exclusions

Our case definition for rLBP excluded patients with serious pathology such as malignancy, fracture, infection and cauda equina syndrome, in accordance with prior, similar studies (online supplemental table 3).^{32 34 55 56} In addition, those with previous lumbar surgery, scoliosis, spondylolisthesis, lumbosacral plexopathy, myelopathy, fibromyalgia and multiple sclerosis were excluded, as these conditions represent alternate causes or mimickers of rLBP^{57 58} and may have a different treatment approach with regards to chiropractic care and gabapentin prescription.

Patients with seizure disorders and epilepsy, diabetic neuropathy, herpetic neuralgia and spinal cord injury

were broadly excluded as these represent FDA-approved indications for gabapentin in the USA.^{59 60} Similarly, patients with restless leg syndrome were excluded as this condition represents an FDA-approved indication for gabapentin enacarbil.⁵⁹ Those with myasthenia gravis and myoclonus, conditions which represent contraindications to gabapentin prescription, were also excluded.¹² All exclusions were made over the year preceding the index date of rLBP diagnosis (figure 1).

Variables

Gabapentin

Gabapentin prescription occurring over a 1-year follow-up window after index rLBP diagnosis was examined using the RxNorm code for gabapentin (25 480). A 1-year follow-up was chosen to account for the natural history of rLBP, which typically improves over a span of 3 months to 1 year.^{61 62} In addition, a 1-year follow-up allowed for comparison to normative data describing the frequency of gabapentin prescription.²¹

As the study design was customised to examine gabapentin alone, it was not possible to examine prescription of other gabapentinoids or anticonvulsants (eg, pregabalin, topiramate). Prescription of pregabalin was factored into our propensity matching model; thus, it could not be recorded as an individual outcome. In addition, similar antiepileptic medications such as pregabalin are less frequently prescribed for LBP compared with gabapentin²¹ and thus may require a larger sample size. Finally, different Controlled Substance Scheduling,⁶³ use indications, precautions and contraindications would require a different study methodology for each medication.

Potential confounders

Propensity score matching was used to reduce bias³⁷ by balancing patient characteristics between the CSMT and usual medical care cohorts which had a known relationship to the outcome of interest, odds of gabapentin prescription (online supplemental table 4).⁶⁴ Key variables present within 365 days of the index diagnosis of rLBP were propensity matched. Covariates with a positive or negative association with gabapentin use or prescription^{63 65–69} or conditions which are common off-label indications for gabapentin^{59 70} were selected for matching based on the available literature:

- ▶ Adjuvant analgesic use (positive):⁶⁶ antiarrhythmics, antidepressants, benzodiazepines, corticosteroids, muscle relaxants, serotonin-norepinephrine reuptake inhibitors, other anticonvulsants (ie, topiramate, pregabalin), tricyclic antidepressants.
- ▶ Anxiety, bipolar disorder and depression (positive).⁶⁷
- ▶ Chronic pain (positive).⁶⁶
- ▶ Demographics: age, sex and race/ethnicity (positive or negative).^{65–67 71}
- ▶ Diabetes (positive).⁶³
- ▶ Emergency department or inpatient visit (negative).⁶⁷
- ▶ Headaches, including migraines (positive).^{59 66 69}

- ▶ Insomnia (positive).⁶⁷
- ▶ Irritable bowel syndrome (positive).^{66 70}
- ▶ Opioid use (positive).^{65 67}
- ▶ Smoking status, current or former (positive).⁶⁵
- ▶ Social determinants: unemployment, problems related to economic circumstances (positive).^{65 66}
- ▶ Substance use disorder (positive).^{63 67}

This study did not exhaustively propensity match for all off-label uses for gabapentin such as interstitial cystitis, hot flashes, hiccups, essential tremors, refractory chronic cough, nausea and vomiting and pruritus.^{70 72}

Evidence suggests that these conditions are either not independently associated with gabapentin use⁶⁸ or are uncommon reasons for prescription.⁶⁹

Study size

A required total sample size of 515 patients was calculated with G*Power (V.3.1.9.7), using a z-test for logistic regression and assuming normal distribution. Parameters included a power of 0.95, two tails, alpha error of 0.05 and OR of 0.67 from a similar study regarding benzodiazepine prescription and CSMT for rLBP.³³ The probability for the alternative hypothesis was 0.20, reflecting the frequency of gabapentin prescription in patients with LBP in a previous study.²¹ This sample appeared feasible given the large CSMT population in our previous similar study also using the TriNetX network.³³

Statistical methods

Key baseline characteristics included in propensity matching were compared using a Pearson χ^2 test for categorical variables and independent-samples t-test for continuous variables. Propensity matching was conducted in real-time using software built into the TriNetX data set viewing platform. Propensity score matching involved 1:1 greedy nearest neighbour matching with a calliper distance of 0.1 pooled SDs of the logit of the propensity score. Odds of gabapentin prescription per cohort were calculated by dividing the number of patients receiving a prescription by the number of patients not receiving a prescription. ORs for gabapentin prescription occurring over a 1-year follow-up were calculated by dividing odds in the CSMT cohort by odds in the cohort receiving usual medical care. We did not perform imputations for missing data.

At the recommendation of peer reviewers, we conducted two post hoc sensitivity analyses. A cumulative incidence graph with 95% CIs was used to illustrate the timing of gabapentin prescription and ascertain if, and when, the incidence curves diverged in relation to the index date of rLBP diagnosis. We also examined the likelihood of a negative control outcome⁴² to provide a marker of residual between-cohort imbalance in patient preference towards receiving pharmacological care. This was described in terms of an OR for receipt of any gastrointestinal medication over the 1-year follow-up window, and was calculated using the same methods described above for gabapentin.

Patient and public involvement

No patient or public involvement.

RESULTS

Participants

Eligible patients were identified from 77 healthcare organisations, 10 of which included CSMT as an offered service. Before propensity matching, there were 1635 patients in the CSMT cohort and 429 778 in the cohort receiving usual medical care. During propensity matching the larger usual medical care cohort diminished in size as patients that did not match were removed, resulting in 1635 patients in each cohort (mean age 36.3±8.6 years, 60% women).

Before matching, there were several between-cohort differences (table 1). For example, the CSMT cohort had a significantly greater percentage of patients who were white and not Hispanic/Latino, and lower representation of other racial and ethnic groups. The CSMT cohort had a greater frequency of 'anxiety, dissociative, stress-related, somatoform and other nonpsychotic mental disorders', mood disorders and prescription of antidepressants, among other differences. After matching, no variables were significantly different between cohorts (ie, $p>0.05$ for each).

Descriptive data

The mean number of data points per patient was high in both cohorts (CSMT 1433; usual medical care 989). After propensity matching, the frequency of several 'unknown' demographic variables was the same in both cohorts: unknown race (19%) unknown sex (1%), unknown age (0%). Unknown ethnicity was similar in both cohorts (14% CSMT, 15% usual medical care). Together, these findings suggested there were inconsequential between-cohort differences regarding missing data. A density graph of propensity scores revealed that cohorts were similar after matching (online supplemental figure 1).

Key results

Gabapentin prescription was less frequent in the CSMT cohort over the 1-year follow-up after rLBP diagnosis both before and after propensity matching. After matching, 4.6% of patients in the CSMT cohort and 8.3% in the usual medical care cohort had received a gabapentin prescription (table 2). After matching, odds of gabapentin prescription over the 1-year follow-up were significantly lower in the CSMT cohort compared with the cohort receiving usual medical care, with an OR (95% CI) of 0.53 (0.40 to 0.71; $p<0.0001$).

Sensitivity analyses

Analysis of the cumulative incidence graph revealed that the incidence of gabapentin prescription was greater in the usual medical care cohort than the CSMT cohort at day 0 (figure 2). The incidence of gabapentin prescription remained higher in the usual medical care cohort for

the duration of follow-up, and the incidence curves and 95% CIs did not overlap at any point during follow-up, suggesting that the incidence was significantly different between cohorts throughout.

After propensity score matching, there was no significant difference in the likelihood of receiving any gastrointestinal medication over 1-year follow-up in the CSMT cohort compared with the usual care cohort (OR 0.89 (0.76–1.04)) with an incidence of 26% (CSMT) and 28% (usual medical care).

DISCUSSION

To our knowledge, this retrospective cohort study was the first to examine the association between CSMT and the likelihood of gabapentin prescription among patients with rLBP and included a large sample size with over 1600 patients per propensity matched cohort. These real-world findings support our hypothesis that adults initially receiving CSMT for rLBP have reduced odds of receiving a gabapentin prescription over a 1-year follow-up period. Our cumulative incidence analysis suggested that much of the difference in likelihood in prescription could be attributed to the care received on the date of diagnosis of rLBP, either being pharmacological (usual medical care) or non-pharmacological (CSMT). Per a negative control outcome, our results were not explained by a patient preference to avoid prescription medications.

In a previous study based on 2018 survey data, 20% of US adults (mean age 52) who visited a medical doctor for LBP over the preceding year were recommended gabapentin.²¹ In comparison, the present study found that only 8% of the usual medical care cohort received a gabapentin prescription. The comparatively lower rate of gabapentin prescription in our study may be due several differences in study design such as: (1) our rigorous selection criteria excluded several conditions positively associated with gabapentin prescription (eg, diabetic neuropathy, restless legs syndrome); (2) our new-user design led to the inclusion of younger patients earlier in their course of care; and (3) our study measured documented prescriptions, rather than recommendation of the medication based on patients' recollection.

Our findings are similar to those of previous studies which demonstrated an association between initial receipt of CSMT and reduced odds of prescription of opioids and benzodiazepines.^{33 35 36} Gabapentin, opioids and benzodiazepines are similarly not recommended by several clinical practice guidelines for acute LBP/rLBP.¹³ Accordingly, our findings add to growing evidence that receipt of CSMT early in the care pathway for new onset LBP/rLBP could lead to greater concordance with these guidelines with respect to medication prescribing practices.^{33 35 36} In addition, our findings are consistent with some authors' recommendations that patients with LBP/rLBP should initiate treatment with non-pharmacological providers such as chiropractors.^{19 73}

**Table 1** Baseline characteristics before and after propensity score matching

Characteristic	Before matching			After matching		
	CSMT	Usual medical care	P value	CSMT	Usual medical care	P value
N	1635	429 778		1635	1635	
Age	36.3±8.6	36.8±8.2	0.0146	36.3±8.6	36.3±8.6	0.8319
Sex						
Female	975 (60%)	245 369 (57%)	0.0383	975 (60%)	977 (60%)	0.9432
Male	658 (40%)	184 328 (43%)	0.0310	658 (40%)	656 (40%)	0.9431
Race						
Black or African American	104 (6%)	67 708 (16%)	<0.0001	104 (6%)	106 (6%)	0.8866
White	1185 (72%)	251 213 (58%)	<0.0001	1185 (72%)	1180 (72%)	0.8451
Asian	24 (1%)	10 149 (2%)	0.0175	24 (1%)	22 (1%)	0.7665
Ethnicity						
Hispanic/Latino	45 (3%)	40 549 (9%)	<0.0001	45 (3%)	56 (3%)	0.2662
Not Hispanic/Latino	1359 (83%)	254 250 (59%)	<0.0001	1359 (83%)	1333 (82%)	0.2333
Conditions						
Anxiety, dissociative, stress-related, somatoform and other non-psychotic mental disorders	329 (20%)	59 655 (14%)	<0.0001	329 (20%)	312 (19%)	0.4539
Mood disorders	213 (13%)	44 974 (10%)	0.0007	213 (13%)	199 (12%)	0.4607
Headache	122 (7%)	24 050 (6%)	0.0011	122 (7%)	119 (7%)	0.8409
Chronic pain, not elsewhere classified	121 (7%)	88 548 (21%)	<0.0001	121 (7%)	108 (7%)	0.3730
Mental and behavioural disorders due to psychoactive substance use	99 (6%)	46 163 (11%)	<0.0001	99 (6%)	92 (6%)	0.6017
Migraine	111 (7%)	22 424 (5%)	0.0044	111 (7%)	93 (6%)	0.1931
Nicotine dependence	75 (5%)	38 047 (9%)	<0.0001	75 (5%)	67 (4%)	0.4925
Insomnia	44 (3%)	9060 (2%)	0.1016	44 (3%)	42 (3%)	0.8270
Diabetes mellitus	41 (3%)	18 164 (4%)	0.0006	41 (3%)	39 (2%)	0.8209
Irritable bowel syndrome	32 (2%)	4920 (1%)	0.0021	32 (2%)	33 (2%)	0.9003
Problems related to employment and unemployment	10 (1%)	514 (<1%)	<0.0001	10 (1%)	10 (1%)	1
Problems related to housing and economic circumstances	10 (1%)	1101 (<1%)	0.0046	10 (1%)	10 (1%)	1
Visits						
Emergency	234 (14%)	147 581 (34%)	<0.0001	234 (14%)	213 (13%)	0.2851
Inpatient	137 (8%)	30 848 (7%)	<0.0604	137 (8%)	125 (7%)	0.4395
Medications						
Medications (any)	1142 (70%)	317 385 (74%)	0.0002	1142 (70%)	1140 (70%)	0.9393
Opioid analgesics	257 (16%)	102 128 (24%)	<0.0001	257 (16%)	244 (15%)	0.5279
Benzodiazepine derivative sedatives/hypnotics	144 (9%)	49 010 (11%)	0.0010	144 (9%)	124 (8%)	0.2023
Antidepressants	326 (20%)	55 737 (13%)	<0.0001	326 (20%)	323 (20%)	0.8954
Antiarrhythmics	75 (5%)	57 859 (13%)	<0.0001	75 (5%)	69 (4%)	0.6091
Glucocorticoids	348 (21%)	132 469 (31%)	<0.0001	348 (21%)	320 (20%)	0.2246
Skeletal muscle relaxants	176 (11%)	117 383 (27%)	0.0446	176 (11%)	175 (11%)	0.9549
Pregabalin	10 (1%)	4107 (1%)	0.1533	10 (1%)	10 (1%)	1
Topiramate	13 (1%)	5582 (1%)	0.0724	13 (1%)	11 (1%)	0.6820

CMST, chiropractic spinal manipulative therapy.

Table 2 Key results before and after propensity score matching

	Before matching		After matching	
	CSMT n=1635	Usual medical care n=4 29 778	CSMT n=1635	Usual medical care n=1635
Gabapentin No. (%)	75 (4.6)	43 314 (9.9)	75 (4.6)	136 (8.3)
OR (95% CI)	0.44 (0.35 to 0.55)*	(Reference)	0.53 (0.40 to 0.71)*	(Reference)

*Indicates a p value of <0.0001.

%, percentage of patients receiving a gabapentin prescription; CSMT, chiropractic spinal manipulative therapy; No., number.

There are several potential explanations as to why initial CSMT for rLBP could be associated with a reduction in gabapentin prescription. First, while US chiropractors are portal-of-entry providers, they do not prescribe medications, including gabapentin. As such, they are not faced with pressure or even the option to prescribe medications for pain. In addition, rLBP generally has a good prognosis, with most patients improving by 1 year.^{61 62} Therefore, we suspect that patients visiting a chiropractor initially for rLBP (1) may improve with CSMT, (2) improve via the favourable natural history of rLBP or (3) enter a non-pharmacological care pathway instead of visiting providers who have medication prescription as part of their scope of practice, and thus be more likely to prescribe gabapentin.

Considering that previous randomised controlled trials have found that CSMT is effective in alleviating LBP⁷⁴ and rLBP,^{75 76} it remains possible that pain relief afforded by CSMT accounts for the observed reduction in gabapentin prescription. However, we are unaware of any studies that examined gabapentin prescription alongside markers of pain and/or disability that could further support this hypothesis. Accordingly, a future pragmatic clinical trial could examine the potential interaction between pain relief and likelihood of gabapentin prescription among patients randomised to enter a chiropractic or medical care pathway for new onset rLBP.

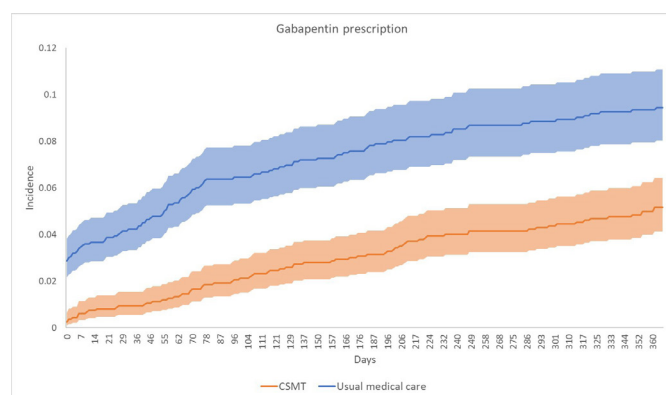


Figure 2 Cumulative incidence graph. Receipt of gabapentin prescription in the chiropractic spinal manipulative therapy cohort (CSMT; orange) versus usual medical care cohort (blue) is illustrated over the 1-year follow-up window (365 days). Shaded regions indicate 95% CIs. CSMT, chiropractic spinal manipulative therapy.

The reduction in absolute risk of gabapentin prescription over 1-year follow-up was relatively small in the present study (4%). However, we cannot rule out a clinically important effect considering the potential risk of abuse, misuse, dependence, withdrawal⁹ and adverse events¹² related to gabapentin use. One previous study found that patients who received CSMT for LBP had significantly reduced odds of having an adverse drug reaction (OR of 0.49),⁷⁷ suggesting that reduced prescription of medications used to treat pain could translate into less adverse events. However, we were unable to examine for the likelihood of potential adverse events related to gabapentin in our study considering: (1) we had limited sample size to evaluate this outcome, (2) our study population was highly selected, via excluding or controlling for comorbid conditions and potential drug interactions and (3) data regarding the dose of gabapentin was unavailable. A follow-up study, if sufficiently powered with a larger data range and data regarding dose, could better examine markers related to clinical significance (eg, adverse events).

Similar retrospective cohort studies should be undertaken to further explore the association between CSMT and gabapentin prescription using other large data sets which may include different patient populations (eg, Medicare, Medicaid) or healthcare settings (eg, Veterans Health Administration, private practices or practice-based research networks). Similar results to the current study would then justify a prospective study, such as a randomised controlled trial, to reduce residual sources of confounding. A prospective trial would also allow for related health outcomes such as changes in health-related quality of life, pain severity, additional social determinants and LBP-related direct or indirect costs to be examined in tandem.

Limitations

This study has several limitations. First, there may be unmeasured confounding. Despite our efforts to control for socioeconomic variables relating to income and education level, these variables may not have been sufficiently represented in the TriNetX data set. Other variables which may influence gabapentin prescription, such as geographical location,⁶⁵ pain severity and LBP-related disability, were unavailable in the data set.

Second, patients could be misclassified. As the study included data derived from medical records, diagnoses or comorbidities could be missing, outdated or incorrect. Metrics regarding data completeness were unavailable for several variables. Our query could not be validated against a gold-standard chart review given that data were de-identified and aggregated from several sources.

Third, patients' eligibility could change during follow-up. For example, patients could have received a diagnosis of diabetic neuropathy after rLBP diagnosis, which was not present at baseline. While this could not be completely prevented, we minimised the potential for between-cohort differences in changing eligibility by the extensive use of propensity score matching at baseline (eg, matching for diabetes mellitus).

Fourth, we were unable to compare gabapentin prescribing rates according to initial provider type as the TriNetX data set does not catalogue provider codes. As rates of gabapentin prescribing may vary across provider type,⁶⁵ this information would allow for a more in-depth analysis. In addition, based on previous literature regarding opioids,^{19 34 36} it is possible that initiating care for rLBP with any non-pharmacological provider (ie, physical therapist, acupuncturist, chiropractor) would similarly yield a reduction in prescribing of gabapentin.

Fifth, this study did not incorporate non-clinical factors such as a pressure to prescribe medications for pain, patients' expectations or providers' concern regarding patient satisfaction surveys, which could influence the likelihood of gabapentin prescription.⁷⁸

Sixth, this study did not examine markers of gabapentin misuse, abuse or illicit use, which are not adequately recorded in the data set. However, our strategy of propensity matching for substance use disorders aimed to minimise confounding related to this possibility.

Seventh, gabapentin prescriptions were temporally but not deterministically linked to rLBP diagnoses; therefore, it remains possible that their prescription may have been for another condition. This possibility was minimised by our strategy to exclude patients with potential on-label gabapentin indications (eg, seizure disorders, diabetic neuropathy), and account for patients with potential off-label uses of gabapentin via exclusion (eg, fibromyalgia)^{59 70} or propensity matching (eg, anxiety, irritable bowel syndrome).^{66 70}

Finally, study results may only be generalisable to large academic healthcare organisations and may not apply to smaller private practice settings. Further, study results may not be generalisable to healthcare settings outside of the USA, which may have varied legal status and guideline recommendations regarding the prescription of gabapentin for rLBP.

CONCLUSION

This large retrospective cohort study found that adults receiving CSMT for a new diagnosis of rLBP have significantly reduced odds of receiving a gabapentin

prescription over 1-year follow-up compared with those receiving usual medical care. According to our sensitivity analyses, the difference in incidence of prescription was largely attributed to the type of care received on the index date of rLBP diagnosis, and was not explained by a preference for patients in the CSMT cohort to avoid prescription medications. These findings are consistent with a potential influence of early CSMT on patients' rLBP care pathway towards avoiding certain prescription medications. However, our findings may not be generalisable to smaller practice settings or other countries and should be replicated and corroborated by a prospective study to reduce residual sources of confounding.

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